

1

## 1.1 Basic probability

**Probability** or **chance** can be measured on a scale which runs from **zero**, which represents **impossibility**, to **one**, which represents **certainty**.

### 1.1.1 Terminology

A **sample space**,  $\Omega$ , is the set of all possible **outcomes** of an experiment. An **event**  $E \in \Omega$  is a subset of  $\Omega$ .

**Example 1** *Experiment:* roll a die twice. Possible *events* are  $E_1 = \{1\text{st face is a } 6\}$ ,  $E_2 = \{\text{sum of faces} = 3\}$ ,  $E_3 = \{\text{sum of faces is odd}\}$ ,  $E_4 = \{1\text{st face} - 2\text{nd face} = 3\}$ . Identify the sample space and the above events. Obtain their probabilities when the die is **fair**.

**Answer:**

		second roll					
first roll	1	(1,1)	(1,2)	(1,3)	(1,4)	(1,5)	(1,6)
	2	(2,1)	(2,2)	(2,3)	(2,4)	(2,5)	(2,6)
	3	(3,1)	(3,2)	(3,3)	(3,4)	(3,5)	(3,6)
	4	(4,1)	(4,2)	(4,3)	(4,4)	(4,5)	(4,6)
	5	(5,1)	(5,2)	(5,3)	(5,4)	(5,5)	(5,6)
	6	(6,1)	(6,2)	(6,3)	(6,4)	(6,5)	(6,6)

$$p(E_1) = \frac{1}{6}; p(E_2) = \frac{1}{18}; p(E_3) = \frac{1}{2}; p(E_4) = \frac{1}{12}.$$

#### Combinations of events

Given events  $A$  and  $B$ , further events can be identified as follows.

- The **complement** of any event  $A$ , written  $\bar{A}$  or  $A^c$ , means that  $A$  does **not** occur.
- The **union** of any two events  $A$  and  $B$ , written  $A \cup B$ , means that  $A$  **or**  $B$  **or** both occur.
- The **intersection** of  $A$  and  $B$ , written as  $A \cap B$ , means that both  $A$  **and**  $B$  **occur**

## 1.1.2 Probability axioms

Let  $\mathcal{F}$  be the class of all events in  $\Omega$ . A **probability (measure)**  $P$  on  $(\Omega, \mathcal{F})$  is a real-valued function satisfying the following three axioms:

1.  $P(E) \geq 0$  for every  $E \in \mathcal{F}$
2.  $P(\Omega) = 1$
3. Suppose the events  $E_1$  and  $E_2$  are mutually exclusive (that is,  $E_1 \cap E_2 = \emptyset$ ).

Then

$$P(E_1 \cup E_2) = P(E_1) + P(E_2)$$

*Some consequences:*

- (i)  $P(\bar{E}) = 1 - P(E)$  (so in particular  $P(\emptyset) = 0$ )
- (ii) For any two events  $E_1$  and  $E_2$  we have the **addition rule**

$$P(E_1 \cup E_2) = P(E_1) + P(E_2) - P(E_1 \cap E_2)$$

**Example 1:** (continued)

Obtain  $P(E_1 \cap E_2)$ ,  $P(E_1 \cup E_2)$ ,  $P(E_1 \cap E_3)$  and  $P(E_1 \cup E_3)$

**Answer:**  $P(E_1 \cap E_2) = P(\emptyset) = 0$

$$P(E_1 \cup E_2) = P(E_1) + P(E_2) = \frac{1}{6} + \frac{1}{18} = \frac{2}{9}$$

$$P(E_1 \cap E_3) = P(6, 1), (6, 3), (6, 5) = \frac{3}{36} = \frac{1}{12}$$

$$P(E_1 \cup E_3) = P(E_1) + P(E_3) - P(E_1 \cap E_3) = \frac{1}{6} + \frac{1}{2} - \frac{1}{12} = \frac{7}{12}$$

[Notes on axioms:

(1) In order to cope with infinite sequences of events, it is necessary to strengthen axiom 3 to

3\*.  $P(\cup_{i=1}^{\infty} E_i) = \sum_{i=1}^{\infty} P(E_i)$  for any sequence  $(E_1, E_2, \dots)$  of mutually exclusive events.

(2) When  $\Omega$  is noncountably infinite, in order to make the theory rigorous it is usually necessary to restrict the class of events  $\mathcal{F}$  to which probabilities are assigned.]

### 1.1.3 Conditional probability

Suppose  $P(E_2) \neq 0$ . The **conditional probability** of the event  $E_1$  given  $E_2$  is defined as

$$P(E_1|E_2) = \frac{P(E_1 \cap E_2)}{P(E_2)}.$$

The conditional probability is undefined if  $P(E_2) = 0$ . The conditional probability formula above yields the **multiplication rule**:

$$\begin{aligned} P(E_1 \cap E_2) &= P(E_1)P(E_2|E_1) \\ &= P(E_2)P(E_1|E_2) \end{aligned}$$

#### Independence

Events  $E_1$  and  $E_2$  are said to be **independent** if

$$P(E_1 \cap E_2) = P(E_1)P(E_2).$$

Note that this implies that  $P(E_1|E_2) = P(E_1)$  and  $P(E_2|E_1) = P(E_2)$ . Thus knowledge of the occurrence of one of the events does not affect the likelihood of occurrence of the other.

Events  $E_1, \dots, E_k$  are **pairwise independent** if  $P(E_i \cap E_j) = P(E_i)P(E_j)$  for all  $i \neq j$ . They are **mutually independent** if for all subsets  $P(\cap_j E_j) = \prod_j P(E_j)$ . Clearly, mutual independence  $\Rightarrow$  pairwise independence, but the converse is false (see question 4 of the self study exercises).

**Example 1** (continued): Find  $P(E_1|E_2)$  and  $P(E_1|E_3)$ . Are  $E_1, E_2$  independent?

**Answer:**  $P(E_1|E_2) = \frac{P(E_1 \cap E_2)}{P(E_2)} = 0$ ,  $P(E_1|E_3) = \frac{P(E_1 \cap E_3)}{P(E_3)} = \frac{1/12}{1/2} = \frac{1}{6}$   
 $P(E_1)P(E_2) \neq 0$  so  $P(E_1 \cap E_2) \neq P(E_1)P(E_2)$  and thus  $E_1$  and  $E_2$  are not independent.

Law of total probability (partition law)

Suppose that  $B_1, \dots, B_k$  are **mutually exclusive** and **exhaustive** events (i.e.  $B_i \cap B_j = \emptyset$  for all  $i \neq j$  and  $\cup_i B_i = \Omega$ ).

Let  $A$  be any event. Then

$$P(A) = \sum_{j=1}^k P(A|B_j)P(B_j)$$

Bayes' Rule

Suppose that events  $B_1, \dots, B_k$  are mutually exclusive and exhaustive and let  $A$  be any event. Then

$$P(B_j|A) = \frac{P(A|B_j)P(B_j)}{P(A)} = \frac{P(A|B_j)P(B_j)}{\sum_i P(A|B_i)P(B_i)}$$

5

**Example 2:** (Cancer diagnosis) A screening programme for a certain type of cancer has reliabilities  $P(A|D) = 0.98$ ,  $P(A|\bar{D}) = 0.05$ , where  $D$  is the event “disease is present” and  $A$  is the event “test gives a positive result”. It is known that 1 in 10, 000 of the population has the disease. Suppose that an individual’s test result is positive. What is the probability that that person has the disease?

**Answer:** We require  $P(D|A)$ . First find  $P(A)$ .

$$P(A) = P(A|D)P(D) + P(A|\bar{D})P(\bar{D}) = 0.98 \times 0.0001 + 0.05 \times 0.9999 = 0.050093.$$

$$\text{By Bayes' rule; } P(D|A) = \frac{P(A|D)P(D)}{P(A)} = \frac{0.0001 \times 0.98}{0.050093} = 0.002.$$

The person is still very unlikely to have the disease even though the test is positive.

**Example 3:** (Bertrand’s Box Paradox) Three indistinguishable boxes contain black and white beads as shown: [ww], [wb], [bb]. A box is chosen at random

and a bead chosen at random from the selected box. What is the probability of that the [wb] box was chosen given that selected bead was white?

**Answer:**  $E \equiv$  ‘chose the [wb] box’,  $W \equiv$  ‘selected bead is white’. By the partition law:  $P(W) = 1 \times \frac{1}{3} + \frac{1}{2} \times \frac{1}{3} + 0 \times \frac{1}{3} = \frac{1}{2}$ . Now using Bayes’ rule  $P(E|W) = \frac{P(E)P(W|E)}{P(W)} = \frac{\frac{1}{3} \times \frac{1}{2}}{\frac{1}{2}} = \frac{1}{3}$  (i.e. even though a bead from the selected box has been seen, the probability that the box is [wb] is still  $\frac{1}{3}$ ).